

*Research Article*

## **Best Possible Sufficient Conditions for Strong Law of Large Numbers for Multi-Indexed Orthogonal Random Elements**

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Received 26 April 2006; Revised 18 December 2006; Accepted 5 February 2007

Recommended by Andrew Rosalsky

It will be shown and induced that the  $d$ -dimensional indices in the Banach spaces version conditions  $\sum_{\mathbf{n}} (E\|X_{\mathbf{n}}\|^p/|\mathbf{n}^\alpha|^p) < \infty$  are sufficient to yield  $\lim_{\min_{1 \leq j \leq d}(n_j) \rightarrow \infty} (1/|\mathbf{n}^\alpha|) \sum_{\mathbf{k} \leq \mathbf{n}} \prod_{j=1}^d (1 - (k_j - 1)/n_j) X_k = 0$  a.s. for arrays of James-type orthogonal random elements. Particularly, it will be shown also that there are the best possible sufficient conditions for multi-indexed independent real-valued random variables.

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### **1. Introduction**

The laws of large numbers for orthogonal random variables or Banach space-valued random elements are investigated by several authors. A consequence of Rademacher-Menshov theorem [2, 3] showed that the sufficient condition for a strong law of large numbers of a sequence of orthogonal real-valued random variables with 0 means and finite second moments is  $\sum_{k=1}^{\infty} (\sigma_k^2/k^2) \cdot [\log_2(k+1)]^2 < \infty$ . Warren and Howell [4] proposed the sufficient condition  $\sum_{k=1}^{\infty} (E\|X_k\|^{1+\alpha}/k^{1+\alpha}) \cdot \log^{1+\alpha} k < \infty$ ,  $0 < \alpha \leq 1$ , for strong convergence of the one-dimensional  $B$ -valued James-type orthogonal random variables. Móricz [5] showed that  $\sum_{i=1}^{\infty} \sum_{k=1}^{\infty} (\sigma_{ik}^2/i^2 k^2) \cdot [\log_2(i+1)]^2 [\log_2(k+1)]^2 < \infty$  is the necessary condition for the strong convergence for arrays of quasi-orthogonal real-valued random variables. Móricz [6] obtained a sufficient condition for strong limit theorems for arrays of quasi-orthogonal real-valued random variables. Móricz et al. [7] showed that the sufficient condition for the strong convergence of  $(1/m^\alpha n^\beta) \sum_{i=1}^m \sum_{k=1}^n X_{ik}$  for arrays of orthogonal, type  $p$  Banach space-valued random elements is

$$\sum_{i=1}^{\infty} \sum_{k=1}^{\infty} \left( \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} \right) \cdot [\log_2(i+1)]^p [\log_2(k+1)]^p < \infty. \quad (1.1)$$

In this paper, the strong laws of large numbers will be investigated for James type of orthogonality in a Banach space. In order to induce  $d$ -dimensional case,  $d > 2$ , we wish to investigate the strongly convergent behavior of a more general Cesaro-type means,  $(1/m^\alpha n^\beta) \sum_{i=1}^m \sum_{k=1}^n (1 - (i-1)/m)(1 - (k-1)/n)X_{ik}$ , as  $m, n \rightarrow \infty$ , for arrays of two-dimensionally indexed orthogonal random elements in a Banach space of type  $p$ ,  $1 \leq p \leq 2$ , and  $1/2 < \alpha, \beta \leq 1$ , though Su [1] showed a case of  $\alpha = 1 = \beta$ . In particular, it will be proven that the sufficient conditions are also the best possible even for independent real-valued random variables. The definition for an array of orthogonal random elements and the formulation of previous results and auxiliary lemmas for orthogonality are given in Section 2. The major results and their proofs are in Sections 3 and 4, respectively.

## 2. Preliminaries and auxiliary lemmas

The basic definitions and properties of Banach space-valued random variables (or random elements) are well established in the literature (e.g., [8]). In these preliminaries, we only introduce the concepts which are necessary and not easy to read in the literature.

Our sense of orthogonality throughout this manuscript is that of *James type orthogonality*. For elements  $x$  and  $y$  in a Banach space  $B$ ,  $x$  is said to be *James orthogonal* to  $y$  (denoted by  $x \perp_J y$ ) if  $\|x\| \leq \|x + ty\|$  for all  $t \in \mathcal{R}$ . If  $B$  is a Hilbert space, then James type orthogonality agrees with the usual notion of orthogonality where the inner product is 0 since  $\|x + ty\|^2 = (x + ty, x + ty) = \|x\|^2 + t^2\|y\|^2 + 2t(x, y) \geq \|x\|^2$  for all  $t \in \mathcal{R}$  if and only if  $(x, y) = 0$ . However, in a Banach space where the norm is not generated by an inner product, it is possible for  $x \perp_J y$  but  $y \not\perp_J x$  with  $(x, y) \neq 0$ . For instance, let  $\mathcal{R}^2 = \{(x_1, x_2) : \|(x_1, x_2)\| = |x_1| + |x_2|, x_1, x_2 \in \mathcal{R}\}$  and  $x = (2, 0)$  and  $y = (2, -2)$ . Then, it is clear that the usual inner product  $(x, y) = 4 \neq 0$ . Next,  $x \perp_J y$  but  $y \not\perp_J x$  since  $\|x + ty\| = |2 + 2t| + |-2t| \geq 2 = \|x\|$  and  $\|y + tx\| = |2 + 2t| + |-2| = 3 < \|y\| = 4$  while picking  $t = -1/2$ . Therefore, it is not possible to create a notation of orthogonality with a good geometrical meaning in an arbitrary Banach space without the inner product. As a result, James-type orthogonality is adopted to circumvent this shortcoming [7].

Let  $\{X_{ik}, i, k \geq 1\}$  be a double sequence of random elements in the Banach space  $L^p(B)$  with zero means, that is,  $E(X_{ik}) = 0$  for all  $i, k$  and finite  $p$ th moments,  $E\|X_{ik}\|^p < \infty$  for all  $i, k$ , where  $\|\cdot\|$  is the norm of the separable Banach space  $B$ . The following is the extended definition for arrays of orthogonal random variables in Banach spaces.

*Definition 2.1.* An array of random elements  $\{X_{ik}\}$  is orthogonal in  $L^p(B)$ ,  $1 \leq p < \infty$ , if

$$(i) \quad E\|X_{ik}\|^p < \infty \quad \forall i, k,$$

$$(ii) \quad E \left\| \sum_{i=1}^{n_1} \sum_{k=1}^{n_2} a_{\pi_1(i), \pi_2(k)} X_{\pi_1(i), \pi_2(k)} \right\|^p \leq E \left\| \sum_{i=1}^{n_1+m_1} \sum_{k=1}^{n_2+m_2} a_{\pi_1(i), \pi_2(k)} X_{\pi_1(i), \pi_2(k)} \right\|^p, \quad (2.1)$$

for all arrays  $\{a_{ik}\} \subseteq \mathcal{R}$ , for all  $n_1, n_2, m_1$ , and  $m_2$ , and for all permutations  $\pi_1, \pi_2$  of the positive integers  $\{1, 2, \dots, m_1 + n_1\}$  and  $\{1, 2, \dots, m_2 + n_2\}$ , respectively.

In retrospect, a separable Banach space  $B$  is of type  $p$  ( $1 \leq p \leq 2$ ) if and only if there is a constant  $C > 0$  (depending on  $B$  only) such that  $E\|\sum_{i=1}^n X_i\|^p \leq C \sum_{i=1}^n E\|X_i\|^p$  when  $\{X_i\}$  are independent random elements with zero means and finite  $p$ th moments [8]. In

order to obtain the desired results, a useful version of moment inequality for arrays that extend the results of Howell and Taylor [9] is listed below.

**PROPOSITION 2.2.** *The following conditions are equivalent:*

- (i)  *$B$  is a Banach space of type  $p$ ,  $1 \leq p \leq 2$ ,*
- (ii) *for each array  $\{X_{mn}\}$  of orthogonal random elements in  $L^p(B)$ , there exists a constant  $C$  (depending on  $\{X_{mn}\}$  and  $B$ ) such that, for all  $m$  and  $n$ ,*

$$E \left\| \sum_{i=1}^m \sum_{k=1}^n X_{ik} \right\|^p \leq C \sum_{i=1}^m \sum_{k=1}^n E \|X_{ik}\|^p. \quad (2.2)$$

The next lemma is from Móricz et al. [7].

**LEMMA 2.3.** *If  $\{X_{ik}; i \geq 1, k \geq 1\}$  is an array of orthogonal (in  $L^p(B)$ ) random elements in a Banach space  $B$  of type  $p$  for some  $1 \leq p \leq 2$ , then*

$$E \left[ \left( \max_{1 \leq j \leq m} \left\| \sum_{i=a+1}^{a+j} \sum_{k=b+1}^{b+n} X_{ik} \right\| \right)^p \right] \leq C_1 (\log_2 2m)^p \sum_{i=a+1}^{a+m} \sum_{k=b+1}^{b+n} E \|X_{ik}\|^p \quad (2.3)$$

for some  $C_1 > 0$  and

$$E \left[ \left( \max_{1 \leq j \leq m} \max_{1 \leq t \leq n} \left\| \sum_{i=a+1}^{a+j} \sum_{k=b+1}^{b+t} X_{ik} \right\| \right)^p \right] \leq C_2 (\log_2 2m)^p (\log_2 2n)^p \sum_{i=a+1}^{a+m} \sum_{k=b+1}^{b+n} E \|X_{ik}\|^p \quad (2.4)$$

for some  $C_2 > 0$ .

The following lemma can be derived directly from Proposition 2.2 and Lemma 2.3. Hence, the proofs are omitted.

**LEMMA 2.4.** *Let  $\{X_{ik}\}$  be an array of orthogonal (in  $L^p(B)$ ) random elements in a Banach space  $B$  of type  $p$  for some  $1 \leq p \leq 2$  and let  $\{a_{ik}\}$  be any array of real numbers, then there exists positive constants  $C_3$  and  $C_4$  such that, for all  $m$  and  $n$ ,  $(a, b \geq 0; m, n \geq 1)$*

$$\begin{aligned} \text{(i)} \quad & E \left\| \sum_{i=a+1}^{a+m} \sum_{k=b+1}^{b+n} a_{ik} X_{ik} \right\|^p \leq C_3 \sum_{i=a+1}^{a+m} \sum_{k=b+1}^{b+n} |a_{ik}|^p E \|X_{ik}\|^p, \\ \text{(ii)} \quad & E \left[ \left( \max_{1 \leq t \leq n} \left\| \sum_{i=a+1}^{a+m} \sum_{k=b+1}^{b+t} a_{ik} X_{ik} \right\| \right)^p \right] \leq C_4 (\log_2 2n)^p \sum_{i=a+1}^{a+m} \sum_{k=b+1}^{b+n} |a_{ik}|^p E \|X_{ik}\|^p. \end{aligned} \quad (2.5)$$

We also need two more crucial lemmas as follows; they are extended from Kronecker's lemma and Shirayev [10], respectively, and will be proven in Section 4.

**LEMMA 2.5** (two-dimensionally indexed version of Kronecker's lemma). *Let  $\{a_m\}$  and  $\{b_n\}$  be sequences of positive increasing numbers, both  $a_m \uparrow \infty$  and  $b_n \uparrow \infty$  when  $m \rightarrow \infty$  and  $n \rightarrow \infty$ , respectively. Let  $\{x_{ij}; i, j \geq 1\}$  be an array of positive numbers such that*

$\sum_{i=1}^{\infty} \sum_{j=1}^{\infty} (x_{ij}/a_i b_j)$  converges. Then,

$$\begin{aligned}
 \text{(i)} \quad & \frac{1}{a_m} \sum_{i=1}^m \sum_{j=1}^n \frac{x_{ij}}{b_j} \rightarrow 0 \quad \text{as } m \rightarrow \infty, \\
 & \frac{1}{b_n} \sum_{i=1}^m \sum_{j=1}^n \frac{x_{ij}}{a_i} \rightarrow 0 \quad \text{as } n \rightarrow \infty, \\
 \text{(ii)} \quad & \frac{1}{a_m b_n} \sum_{i=1}^m \sum_{j=1}^n x_{ij} \rightarrow 0 \quad \text{as } m \rightarrow \infty \text{ or } n \rightarrow \infty.
 \end{aligned} \tag{2.6}$$

LEMMA 2.6. A sufficient and necessary condition that  $\zeta_{mn} \rightarrow 0$  with probability one as  $m, n \rightarrow \infty$  is that for any  $\varepsilon > 0$ ,

$$P\left(\sup_{i \geq m} \sup_{k \geq n} \|\zeta_{ik}\| > \varepsilon\right) \rightarrow 0 \quad \text{as } m \rightarrow \infty, n \rightarrow \infty. \tag{2.7}$$

### 3. Major results

Theorems 3.1 and 3.2 are two-dimensionally indexed versions of strong convergence for Cesaro-type means for arrays of Banach space-valued random elements and hence their proofs are more complicated than that in real cases because of the structure of spaces.

THEOREM 3.1. Let  $\{X_{ik}\}$  be an array of orthogonal (in  $L^p(B)$ ) random elements with zero means in a Banach space  $B$  of type  $p$ , for some  $1 \leq p \leq 2$ . If  $\sum_{i=1}^{\infty} \sum_{k=1}^{\infty} (E\|X_{ik}\|^p / i^{\alpha p} k^{\beta p}) < \infty$ ,  $1/2 < \alpha, \beta \leq 1$ . Then,

$$\lim_{m,n \rightarrow \infty} \left\| \frac{1}{m^{\alpha} n^{\beta}} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) \left(1 - \frac{k-1}{n}\right) X_{ik} \right\| = 0 \quad \text{a.s.} \tag{3.1}$$

THEOREM 3.2. Let  $\{X_{ik}\}$  be an array of orthogonal (in  $L^p(B)$ ) random elements with zero means in a Banach space  $B$  of type  $p$ , for some  $1 \leq p \leq 2$ . If  $\sum_{i=1}^{\infty} \sum_{k=1}^{\infty} (E\|X_{ik}\|^p / i^{\alpha p} k^{\beta p}) \cdot [\log_2(k+1)]^p < \infty$ ,  $1/2 < \alpha, \beta \leq 1$ . Then,

$$\lim_{m,n \rightarrow \infty} \left\| \frac{1}{m^{\alpha} n^{\beta}} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) X_{ik} \right\| = 0 \quad \text{a.s.} \tag{3.2}$$

The generalization to  $d$ -dimensional arrays random elements of the previous two theorems can be obtained easily by the same methods [1]. Theorems 3.3 and 3.4 are to show that the sufficient conditions in the previous theorems are the best possible conditions for independent real-valued random variables, since the real line is of type  $p$ ,  $1 \leq p \leq 2$ .

THEOREM 3.3. If  $\{\tau_{ik}\}$  is an array of nonnegative real numbers such that  $\tau_{ik}^p / i^{\alpha(p-1)} k^{\beta(p-1)} \leq 1$  holds for indices  $i$  and  $k$  and the condition  $\sum_{i=1}^{\infty} \sum_{k=1}^{\infty} (\tau_{ik}^p / i^{\alpha p} k^{\beta p}) = \infty$ , for some  $1 \leq p \leq 2$  and  $1/2 < \alpha, \beta \leq 1$ , then there exists an array  $\{X_{ik}\}$  of independent real-valued random

variables such that  $E(X_{ik}) = 0$ ,  $E|X_{ik}|^p = \tau_{ik}^p$ , and

$$\limsup_{m+n \rightarrow \infty} \left| \frac{1}{m^\alpha n^\beta} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) \left(1 - \frac{k-1}{n}\right) X_{ik} \right| = \infty. \quad (3.3)$$

**THEOREM 3.4.** *If  $\{\tau_{ik}\}$  is an array of nonnegative real numbers such that  $\tau_{ik}^p \cdot \log_2^{p-1}(k+1)/i^{\alpha(p-1)}k^{\beta(p-1)} \leq 1$  holds for indices  $i$  and  $k$  and the condition  $\sum_{i=1}^{\infty} \sum_{k=1}^{\infty} (\tau_{ik}^p / i^{\alpha p} k^{\beta p}) \cdot [\log_2(k+1)]^p = \infty$ , for some  $1 \leq p \leq 2$  and  $1/2 < \alpha, \beta \leq 1$ , then there exists an array  $\{X_{ik}\}$  of independent real-valued random variables such that  $E(X_{ik}) = 0$ ,  $E|X_{ik}|^p = \tau_{ik}^p$ , and*

$$\limsup_{m+n \rightarrow \infty} \left| \frac{1}{m^\alpha n^\beta} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) X_{ik} \right| = \infty. \quad (3.4)$$

#### 4. Proofs

Here we will verify Lemmas 2.5 and 2.6 first, then prove the case of  $d = 2$ , that is, Theorems 3.1, 3.2, 3.3, and 3.4. We may apply the analogous approaches for the  $d$ -dimensional cases,  $d > 2$ .

*Proof of Lemma 2.5.* (i) First, we have  $\sum_{i=1}^{\infty} (x_{ij}/a_i b_j) < \infty$  for each  $j \geq 1$  and  $\sum_{j=1}^{\infty} (x_{ij}/a_i b_j) < \infty$  for each  $i \geq 1$ . Then by the one-dimensional version Kronecker's lemma, we can conclude that  $(1/a_m) \sum_{i=1}^m (x_{ij}/b_j) \rightarrow 0$  as  $m \rightarrow \infty$ , for every  $j$ , and  $(1/b_n) \sum_{j=1}^n (x_{ij}/a_i) \rightarrow 0$  as  $n \rightarrow \infty$ , for every  $i$ . Hence, for any  $\varepsilon_1 > 0$ , choose  $n > N$  such that  $(1/b_n) \sum_{j=1}^n (x_{ij}/a_i) \leq \varepsilon_1/2^i$  for all  $i$ . Then, we have

$$\frac{1}{b_n} \sum_{i=1}^m \sum_{j=1}^n \frac{x_{ij}}{a_i} = \sum_{i=1}^m \frac{1}{b_n} \sum_{j=1}^n \frac{x_{ij}}{a_i} \leq \sum_{i=1}^m \frac{\varepsilon_1}{2^i} < \varepsilon_1. \quad (4.1)$$

Similarly, for any  $\varepsilon_2 > 0$ , choose  $m > M$  such that  $(1/a_m) \sum_{j=1}^m (x_{ij}/b_j) \leq \varepsilon_2/2^j$  for all  $j$ , then

$$\frac{1}{a_m} \sum_{i=1}^m \sum_{j=1}^n \frac{x_{ij}}{b_j} = \sum_{j=1}^n \frac{1}{a_m} \sum_{i=1}^m \frac{x_{ij}}{b_j} \leq \sum_{j=1}^n \frac{\varepsilon_2}{2^j} < \varepsilon_2. \quad (4.2)$$

(ii) Apparently, for any  $\varepsilon > 0$ , when  $m > M$  (say), we can conclude that

$$\frac{1}{a_m b_n} \sum_{i=1}^m \sum_{j=1}^n x_{ij} \leq \frac{1}{a_m} \sum_{i=1}^m \sum_{j=1}^n \frac{x_{ij}}{b_j} < \varepsilon. \quad (4.3)$$

□

*Proof of Lemma 2.6.* Fix any  $m \geq 1$ , for any  $\varepsilon > 0$ , let  $A_{mn}^\varepsilon = \{\omega : \sup_{i \geq m} \|\zeta_{in}\| > \varepsilon\}$  and  $A_m^\varepsilon = \overline{\bigcap_{n=1}^\infty \bigcup_{k \geq n} A_{mk}^\varepsilon}$ . Then,  $\{\omega : \sup_{i \geq m} \|\zeta_{in}\| \not\rightarrow 0, \text{ as } n \rightarrow \infty\} = \bigcup_{\varepsilon \geq 0} A_m^\varepsilon = \bigcup_{t=1}^\infty A_m^{1/t}$ . However,  $P(A_m^\varepsilon) = \lim_n P(\bigcup_{k \geq n} A_{mk}^\varepsilon)$ . Hence, for any fixed  $m \geq 1$ ,

$$\begin{aligned}
 0 &= P\left(\omega : \sup_{i \geq m} \|\zeta_{in}\| \not\rightarrow 0 \text{ as } n \rightarrow \infty\right) = P\left(\bigcup_{\varepsilon \geq 0} A_m^\varepsilon\right) \\
 &\iff P\left(\bigcup_{t=1}^\infty A_m^{1/t}\right) = 0 \quad \text{any fixed } m \geq 1 \\
 &\iff P(A_m^{1/t}) = 0, \quad t \geq 1, \text{ any fixed } m \geq 1 \\
 &\iff P(A_m^\varepsilon) = 0 \quad \text{for any } \varepsilon > 0, \text{ any fixed } m \geq 1 \\
 &\iff P\left(\bigcup_{k \geq n} A_{mk}^\varepsilon\right) \rightarrow 0 \quad \text{as } n \rightarrow \infty, \text{ any fixed } m \geq 1 \\
 &\iff P\left(\sup_{k \geq n} \sup_{i \geq m} \|\zeta_{ik}\| > \varepsilon\right) \rightarrow 0 \quad \text{as } n \rightarrow \infty, \text{ any fixed } m \geq 1 \\
 &\iff P\left(\bigcap_{m=1}^\infty \left\{ \sup_{k \geq n} \sup_{i \geq m} \|\zeta_{ik}\| > \varepsilon \right\}\right) \rightarrow 0 \quad \text{as } n \rightarrow \infty \\
 &\iff P\left(\sup_{k \geq n} \sup_{i \geq m} \|\zeta_{ik}\| > \varepsilon\right) \rightarrow 0 \quad \text{as } m \rightarrow \infty, n \rightarrow \infty.
 \end{aligned} \tag{4.4}$$

□

*Proof of Theorem 3.1.* We need some useful arguments in the proof in [1, 7], and Lemmas 2.3 and 2.4. For positive integers  $u$  and  $v$ , for any  $\varepsilon$ ,

$$P\left[\sup_{m > 2^u, n > 2^v} \|\xi_{mn}\| > \varepsilon\right] \leq \sum_{r=u}^\infty \sum_{s=v}^\infty P\left[\max_{2^r < m \leq 2^{r+1}} \max_{2^s < n \leq 2^{s+1}} \|\xi_{mn}\| > \varepsilon\right], \tag{4.5}$$

where

$$\xi_{mn} = \frac{1}{m^\alpha n^\beta} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) \left(1 - \frac{k-1}{n}\right) X_{ik}. \tag{4.6}$$

We have

$$\max_{2^r < m \leq 2^{r+1}} \max_{2^s < n \leq 2^{s+1}} \|\xi_{mn}\| \leq \|\xi_{2^r, 2^s}\| + \sum_{j=1}^3 A_{rs}^{(j)}, \tag{4.7}$$

where

$$\begin{aligned}
 A_{rs}^{(1)} &= \max_{2^r < m \leq 2^{r+1}} \|\xi_{m, 2^s} - \xi_{2^r, 2^s}\|, \\
 A_{rs}^{(2)} &= \max_{2^s < n \leq 2^{s+1}} \|\xi_{2^r, n} - \xi_{2^r, 2^s}\|, \\
 A_{rs}^{(3)} &= \max_{2^r < m \leq 2^{r+1}} \max_{2^s < n \leq 2^{s+1}} \|\xi_{mn} - \xi_{m, 2^s} - \xi_{2^r, n} + \xi_{2^r, 2^s}\|.
 \end{aligned} \tag{4.8}$$

Therefore,

$$P\left[\max_{2^r \leq m \leq 2^{r+1}} \max_{2^s \leq n \leq 2^{s+1}} \|\xi_{mn}\| > \varepsilon\right] \leq P\left[\|\xi_{2^r, 2^s}\| > \frac{\varepsilon}{4}\right] + \sum_{j=1}^3 P\left[A_{rs}^{(j)} > \frac{\varepsilon}{4}\right]. \quad (4.9)$$

First, by Markov's inequality and Lemmas 2.3 and 2.4,

$$\begin{aligned} & \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P\left[\|\xi_{2^r, 2^s}\| > \frac{\varepsilon}{4}\right] \\ & \leq \left(\frac{4}{\varepsilon}\right)^p \Gamma_1 \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \frac{1}{2^{\alpha r p} 2^{\beta s p}} \sum_{i=1}^{2^r} \sum_{k=1}^{2^s} \left|1 - \frac{i-1}{2^r}\right|^p \left|1 - \frac{k-1}{2^s}\right|^p E\|X_{ik}\|^p \\ & \leq \left(\frac{4}{\varepsilon}\right)^p \Gamma_1 \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \frac{1}{2^{\alpha r p} 2^{\beta s p}} \left\{ \sum_{i=1}^{2^u} \sum_{k=1}^{2^v} + \sum_{i=2^u+1}^{2^r} \sum_{k=1}^{2^v} + \sum_{i=1}^{2^u} \sum_{k=2^v+1}^{2^s} + \sum_{i=2^u+1}^{2^r} \sum_{k=2^v+1}^{2^s} \right\} E\|X_{ik}\|^p \\ & = \left(\frac{4}{\varepsilon}\right)^p \Gamma_1 \sum_{j=1}^4 B_{uv}^{(j)}, \quad \text{say, for some } \Gamma_1 > 0. \end{aligned} \quad (4.10)$$

Using some basic calculation, it follows that

$$B_{uv}^{(1)} = \frac{2^{(\alpha+\beta)p}}{(2^{\alpha p} - 1)(2^{\beta p} - 1)} \cdot \frac{1}{2^{\alpha p u} 2^{\beta p v}} \sum_{i=1}^{2^u} \sum_{k=1}^{2^v} E\|X_{ik}\|^p. \quad (4.11)$$

Next,

$$B_{uv}^{(2)} \leq \frac{2^{(\alpha+\beta)p}}{(2^{\alpha p} - 1)(2^{\beta p} - 1)} \left\{ \sum_{i=2^u+1}^{\infty} \sum_{k=1}^{2^v} \frac{1}{2^{\beta p v}} \frac{E\|X_{ik}\|^p}{i^{\alpha p}} + \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} \right\}. \quad (4.12)$$

Similarly,

$$\begin{aligned} B_{uv}^{(3)} & \leq \frac{2^{(\alpha+\beta)p}}{(2^{\alpha p} - 1)(2^{\beta p} - 1)} \left\{ \sum_{i=1}^{2^u} \sum_{k=2^v+1}^{\infty} \frac{1}{2^{\alpha p u}} \frac{E\|X_{ik}\|^p}{k^{\beta p}} + \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} \right\}, \\ B_{uv}^{(4)} & = \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \frac{1}{2^{\alpha r p} 2^{\beta s p}} \sum_{i=2^u+1}^{2^r} \sum_{k=2^v+1}^{2^s} E\|X_{ik}\|^p \leq \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}}. \end{aligned} \quad (4.13)$$

Secondly, since

$$A_{rs}^{(1)} = \max_{2^r < m \leq 2^{r+1}} \|\xi_{m, 2^s} - \xi_{2^r, 2^s}\| = \max_{1 \leq m \leq 2^r} \left\| \sum_{t=2^r+1}^{2^{r+m}} (\xi_{t, 2^s} - \xi_{t-1, 2^s}) \right\|, \quad (4.14)$$

where

$$\xi_{t,2^s} - \xi_{t-1,2^s} = \sum_{i=1}^t \sum_{k=1}^{2^s} a_{ik}(t,s) X_{ik}, \quad (4.15)$$

$$a_{ik}(t,s) = \frac{1}{2^{\beta s}} \left(1 - \frac{k-1}{2^s}\right) \left[ (i-1) \left( \frac{1}{(t-1)^{\alpha+1}} - \frac{1}{t^{\alpha+1}} \right) - \left( \frac{1}{(t-1)^\alpha} - \frac{1}{t^\alpha} \right) \right],$$

we have

$$|a_{ik}(t,s)| \leq \frac{1}{t^{\alpha+1} 2^{\beta s}}. \quad (4.16)$$

Hence, for some  $\Gamma_2, \Gamma_2^* > 0$ ,

$$\begin{aligned} & \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ A_{rs}^{(1)} > \frac{\varepsilon}{4} \right] \\ & \leq \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ \max_{1 \leq m \leq 2^r} \left\| \sum_{t=2^r+1}^{2^r+m} (\xi_{t,2^s} - \xi_{t-1,2^s}) \right\| > \frac{\varepsilon}{4} \right] \\ & \leq \left( \frac{4}{\varepsilon} \right)^p \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} E \left[ \left( \max_{1 \leq m \leq 2^r} \left\| \sum_{t=2^r+1}^{2^r+m} (\xi_{t,2^s} - \xi_{t-1,2^s}) \right\| \right)^p \right] \\ & \leq \left( \frac{4}{\varepsilon} \right)^p \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \Gamma_2 [\log_2 2 \cdot 2^r]^p \sum_{m=2^r+1}^{2^{r+1}} E \|\xi_{m,2^s} - \xi_{m-1,2^s}\|^p \\ & \leq \left( \frac{4}{\varepsilon} \right)^p \Gamma_2^* \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \frac{(r+1)^p}{2^{s\beta p}} \sum_{m=2^r+1}^{2^{r+1}} \sum_{i=1}^m \sum_{k=1}^{2^s} \frac{E \|X_{ik}\|^p}{m^{(\alpha+1)p}} \\ & \approx \left( \frac{4}{\varepsilon} \right)^p \Gamma_2^* \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \sum_{m=2^r+1}^{2^{r+1}} \sum_{i=1}^m \sum_{k=1}^{2^s} \frac{r^p E \|X_{ik}\|^p}{m^{(\alpha+1)p} 2^{\beta s p}} \\ & \leq \left( \frac{4}{\varepsilon} \right)^p \Gamma_2^* \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \sum_{k=1}^{2^s} \frac{r^p}{2^{\beta s p}} \left[ \sum_{i=1}^{2^r} \sum_{m=2^r+1}^{2^{r+1}} + \sum_{i=2^r+1}^{2^{r+1}} \sum_{m=i}^{2^{r+1}} \right] \frac{E \|X_{ik}\|^p}{m^{(\alpha+1)p}} \\ & \leq \left( \frac{4}{\varepsilon} \right)^p \Gamma_2^* \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \sum_{k=1}^{2^s} \frac{r^p}{2^{\beta s p}} \left[ \sum_{i=1}^{2^r} \frac{1}{2^{r[(\alpha+1)p-1]}} + \sum_{i=2^r+1}^{2^{r+1}} \frac{1}{i^{(\alpha+1)p-1}} \right] E \|X_{ik}\|^p \\ & \leq \left( \frac{4}{\varepsilon} \right)^p \Gamma_2^* \sum_{s=v}^{\infty} \sum_{k=1}^{2^s} \frac{1}{2^{\beta s p}} \left\{ \sum_{r=u}^{\infty} \sum_{i=1}^{2^r} \frac{r^p}{2^{r[(\alpha+1)p-1]}} + \sum_{r=u}^{\infty} \sum_{i=2^r+1}^{2^{r+1}} \frac{r^p}{i^{(\alpha+1)p-1}} \right\} E \|X_{ik}\|^p. \end{aligned} \quad (4.17)$$

Since the first term and the second term in the bracket of (4.17) can be expressed, respectively, as

$$\begin{aligned} \sum_{r=u}^{\infty} \sum_{i=1}^{2^r} \frac{r^p}{2^{r[(\alpha+1)p-1]}} &\leq \sum_{i=1}^{\infty} \sum_{u < r; i < 2^r} \frac{r^2}{2^{r[(\alpha+1)p-1]}} \leq 2^{(\alpha+1)p} \left( \sum_{i=1}^{2^u} \frac{1}{i^{(\alpha+1)p-1}} + \sum_{i=2^u+1}^{\infty} \frac{1}{i^{(\alpha+1)p-1}} \right), \\ \sum_{r=u}^{\infty} \sum_{i=2^{r+1}}^{2^{r+1}} \frac{r^p}{i^{(\alpha+1)p-1}} &\leq \sum_{i=2^{u+1}}^{\infty} \sum_{u < r; i < 2^{r+1}} \frac{r^2}{2^{r[(\alpha+1)p-1]}} \leq 2^{(\alpha+1)p} \sum_{i=2^{u+1}}^{\infty} \frac{1}{i^{(\alpha+1)p-1}}, \end{aligned} \quad (4.18)$$

hence,

$$\begin{aligned} (4.17) &\leq 2^{(\alpha+1)p} \left( \frac{4}{\varepsilon} \right)^p \Gamma_2^* \sum_{s=v}^{\infty} \left[ \sum_{k=1}^{2^v} + \sum_{k=2^{v+1}}^{2^s} \right] \left[ \sum_{i=1}^{2^u} + 2 \sum_{i=2^u+1}^{\infty} \right] \frac{E\|X_{ik}\|^p}{i^{\alpha p} 2^{\beta s p}} \\ &\leq 2^{\alpha p} \left( \frac{8}{\varepsilon} \right)^p \Gamma_2^* \left\{ \frac{1}{2^{\beta v p}} \sum_{k=1}^{2^v} \sum_{i=1}^{2^u} \frac{E\|X_{ik}\|^p}{i^{\alpha p}} + \frac{1}{2^{\beta v p}} \sum_{k=1}^{2^v} \sum_{i=2^u+1}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p}} \right. \\ &\quad \left. + \sum_{k=2^{v+1}}^{\infty} \sum_{i=1}^{2^u} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} + 2 \sum_{k=2^{v+1}}^{\infty} \sum_{i=2^u+1}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} \right\}. \end{aligned} \quad (4.19)$$

Similarly, for some  $\Gamma_3 > 0$ ,

$$\begin{aligned} \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ A_{rs}^{(2)} > \frac{\varepsilon}{4} \right] &\leq 2^{\beta p} \left( \frac{8}{\varepsilon} \right)^p \Gamma_3 \left\{ \frac{1}{2^{\alpha u p}} \sum_{k=1}^{2^v} \sum_{i=1}^{2^u} \frac{E\|X_{ik}\|^p}{k^{\beta p}} + \frac{1}{2^{\alpha u p}} \sum_{i=1}^{2^u} \sum_{k=2^{v+1}}^{\infty} \frac{E\|X_{ik}\|^p}{k^{\beta p}} \right. \\ &\quad \left. + \sum_{i=2^{u+1}}^{\infty} \sum_{k=1}^{2^v} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} + 2 \sum_{i=2^{u+1}}^{\infty} \sum_{k=2^{v+1}}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} \right\}. \end{aligned} \quad (4.20)$$

Next, since

$$\begin{aligned} \max_{2^r < m \leq 2^{r+1}} \max_{2^s < n \leq 2^{s+1}} &\| \xi_{mn} - \xi_{m,2^s} - \xi_{2^r,n} + \xi_{2^r,2^s} \| \\ &\leq \max_{1 \leq m \leq 2^r} \max_{1 \leq n \leq 2^s} \left\| \sum_{t_1=2^r+1}^{2^{r+m}} \sum_{t_2=2^s+1}^{2^{s+n}} (\xi_{t_1,t_2} - \xi_{(t_1-1),t_2} - \xi_{t_1,(t_2-1)} + \xi_{(t_1-1),(t_2-1)}) \right\|, \\ \xi_{t_1,t_2} - \xi_{(t_1-1),t_2} - \xi_{t_1,(t_2-1)} + \xi_{(t_1-1),(t_2-1)} &= \sum_{i=1}^{t_1} \sum_{k=1}^{t_2} b_{ik}(t_1, t_2) X_{ik}, \end{aligned} \quad (4.21)$$

where

$$\begin{aligned} b_{ik}(t_1, t_2) &= \left[ \frac{1}{t_1^{\alpha}} \left( 1 - \frac{i-1}{t_1} \right) - \frac{1}{(t_1-1)^{\alpha}} \left( 1 - \frac{i-1}{t_1-1} \right) \right] \\ &\quad \cdot \left[ \frac{1}{t_2^{\beta}} \left( 1 - \frac{k-1}{t_2} \right) - \frac{1}{(t_2-1)^{\beta}} \left( 1 - \frac{k-1}{t_2-1} \right) \right], \end{aligned} \quad (4.22)$$

so apparently,

$$|b_{ik}(t_1, t_2)| \leq \frac{(i-1)(k-1)}{t_1^{\alpha+1} t_2^{\beta+1}} \cdot \frac{1}{(t_1-1)(t_2-1)} \leq \frac{1}{t_1^{\alpha+1} t_2^{\beta+1}}. \quad (4.23)$$

Hence, for some  $\Gamma_4 > 0$ ,

$$\begin{aligned} & \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ A_{rs}^{(3)} > \frac{\varepsilon}{4} \right] \\ & \leq \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ \max_{1 \leq m \leq 2^r} \max_{1 \leq n \leq 2^s} \left\| \sum_{t_1=2^r+1}^{2^r+m} \sum_{t_2=2^s+1}^{2^s+n} (\xi_{t_1, t_2} - \xi_{(t_1-1), t_2} - \xi_{t_1, (t_2-1)} + \xi_{(t_1-1), (t_2-1)}) \right\| > \frac{\varepsilon}{4} \right] \\ & \leq \left( \frac{4}{\varepsilon} \right)^p \Gamma_4 \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} (\log_2 2 \cdot 2^r)^p (\log_2 2 \cdot 2^s)^p \sum_{m=2^r+1}^{2^{r+1}} \sum_{n=2^s+1}^{2^{s+1}} \sum_{i=1}^m \sum_{k=1}^n \frac{E\|X_{ik}\|^p}{m^{(\alpha+1)p} n^{(\beta+1)p}} \\ & \leq \left( \frac{4}{\varepsilon} \right)^p \Gamma_4 \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} (r+1)^p (s+1)^p \sum_{m=2^r+1}^{2^{r+1}} \sum_{n=2^s+1}^{2^{s+1}} \sum_{i=1}^m \sum_{k=1}^n \frac{E\|X_{ik}\|^p}{m^{(\alpha+1)p} n^{(\beta+1)p}} \\ & \approx \left( \frac{4}{\varepsilon} \right)^p \Gamma_4 \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} r^p s^p \left( \sum_{i=1}^{2^r} \sum_{m=2^r+1}^{2^{r+1}} + \sum_{i=2^r+1}^{2^{r+1}} \sum_{m=i}^{2^{r+1}} \right) \left( \sum_{k=1}^{2^s} \sum_{n=2^s+1}^{2^{s+1}} + \sum_{k=2^s+1}^{2^{s+1}} \sum_{n=i}^{2^{s+1}} \right) \frac{E\|X_{ik}\|^p}{m^{(\alpha+1)p} n^{(\beta+1)p}} \\ & \leq 2^{(\alpha+\beta)p} \left( \frac{16}{\varepsilon} \right)^p \Gamma_4 \left\{ \frac{1}{2^{\alpha u p} 2^{\beta v p}} \sum_{i=1}^{2^u} \sum_{k=1}^{2^v} E\|X_{ik}\|^p + \frac{2}{2^{\alpha u p}} \sum_{i=1}^{2^u} \sum_{k=2^v+1}^{\infty} \frac{E\|X_{ik}\|^p}{k^{\beta p}} \right. \\ & \quad \left. + \frac{2}{2^{\beta v p}} \sum_{i=2^u+1}^{\infty} \sum_{k=1}^{2^v} \frac{E\|X_{ik}\|^p}{i^{\alpha p}} + 4 \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} \right\}. \end{aligned} \quad (4.24)$$

Combining the results in (4.9), (4.10), (4.19), (4.20), and (4.24), we can conclude that  $P[\lim_{m,n \rightarrow \infty} \|\xi_{mn}\| = 0] = 1$ . Since  $P[\sup_{m \geq s, n \geq t} \|\xi_{mn}\| > \varepsilon] \rightarrow 0$  as  $s, t \rightarrow \infty$  by applying Lemmas 2.5 and 2.6 on the right-hand side of the previous inequalities, then the proof is completed by a convention that  $\sum_{i=1}^{2^r} \sum_{k=1}^{2^r} = 0 = \sum_{i=1}^{\infty} \sum_{k=1}^{2^r} = \sum_{i=1}^{2^r} \sum_{k=1}^{\infty}$  and  $\sum_{i=2^r}^{\infty} \sum_{k=2^r}^{\infty} = \sum_{i=1}^{\infty} \sum_{k=1}^{\infty}$  if  $r = -1$ .  $\square$

*Proof of Theorem 3.2.* Similar to the previous proof, we start out from assuming that  $2^r \leq m \leq 2^{r+1}$  and  $2^s \leq n \leq 2^{s+1}$  with nonnegative integers  $u$  and  $v$ , and letting

$$\tau_{mn} = \frac{1}{m^{\alpha} n^{\beta}} \sum_{i=1}^m \sum_{k=1}^n \left( 1 - \frac{i-1}{m} \right) X_{ik}. \quad (4.25)$$

First, we can write

$$\sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ \max_{2^r \leq m \leq 2^{r+1}} \max_{2^s \leq n \leq 2^{s+1}} \|\tau_{mn}\| > \varepsilon \right] \leq \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ \|\tau_{2^r, 2^s}\| > \frac{\varepsilon}{4} \right] + \sum_{j=1}^3 D_{rs}^{(j)}, \quad (4.26)$$

where

$$\begin{aligned} D_{rs}^{(1)} &= \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ \max_{2^r < m \leq 2^{r+1}} \|\tau_{m, 2^s} - \tau_{2^r, 2^s}\| > \frac{\varepsilon}{4} \right], \\ D_{rs}^{(2)} &= \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ \max_{2^s < n \leq 2^{s+1}} \|\tau_{2^r, n} - \tau_{2^r, 2^s}\| > \frac{\varepsilon}{4} \right], \\ D_{rs}^{(3)} &= \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ \max_{2^r < m \leq 2^{r+1}} \max_{2^s < n \leq 2^{s+1}} \|\tau_{mn} - \tau_{m, 2^s} - \tau_{2^r, n} + \tau_{2^r, 2^s}\| > \frac{\varepsilon}{4} \right]. \end{aligned} \quad (4.27)$$

Comparing (4.9) to (4.24), we can easily obtain that for some  $\Gamma_5 > 0$ ,

$$\sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ \|\tau_{2^r, 2^s}\| > \frac{\varepsilon}{4} \right] \leq \left( \frac{4}{\varepsilon} \right)^p \Gamma_5 \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} \quad (4.28)$$

and for some  $\Gamma_6, \Gamma_6^* > 0$ ,

$$\begin{aligned} D_{rs}^{(1)} &\leq \left( \frac{4}{\varepsilon} \right)^p \Gamma_6 \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} [\log_2 2 \cdot 2^r]^p \sum_{m=2^r+1}^{2^{r+1}} \sum_{i=1}^m \sum_{k=1}^{2^s} \frac{E\|X_{ik}\|^p}{m^{(\alpha+1)p} 2^{\beta sp}} \\ &\leq 2^{\alpha p} \left( \frac{8}{\varepsilon} \right)^p \Gamma_6^* \left\{ \frac{1}{2^{\beta vp}} \sum_{i=1}^{2^u} \sum_{k=1}^{2^v} \frac{E\|X_{ik}\|^p}{i^{\alpha p}} + \frac{1}{2^{\beta vp}} \sum_{i=2^u+1}^{\infty} \sum_{k=1}^{2^v} \frac{E\|X_{ik}\|^p}{i^{\alpha p}} \right. \\ &\quad \left. + \sum_{i=1}^{2^u} \sum_{k=2^v+1}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} + 2 \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E\|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} \right\}. \end{aligned} \quad (4.29)$$

Next, we need to examine that

$$\begin{aligned} \max_{2^s < n \leq 2^{s+1}} \|\tau_{2^r, n} - \tau_{2^r, 2^s}\| &= \max_{2^s < n \leq 2^{s+1}} \left\| \sum_{t=2^s+1}^{2^{s+1}} (\tau_{2^r, t} - \tau_{2^r, t-1}) \right\|, \\ \tau_{2^r, t} - \tau_{2^r, t-1} &= \sum_{i=1}^{2^r} \sum_{k=1}^t c_{ik}(r, t) X_{ik}, \end{aligned} \quad (4.30)$$

where

$$c_{ik}(r, t) = \begin{cases} \frac{1}{2^{\alpha r}} \left( 1 - \frac{i-1}{2^r} \right) \left[ \frac{1}{t^{\beta}} - \frac{1}{(t-1)^{\beta}} \right], & k = 1, 2, \dots, t-1, \\ \frac{1}{2^{\alpha r}} \left( 1 - \frac{i-1}{2^r} \right) \cdot \frac{1}{t^{\beta}}, & k = t. \end{cases} \quad (4.31)$$

Hence, we use basic calculus to obtain

$$|c_{ik}(r, t)| \leq \frac{\beta}{2^{\alpha r} t^{\beta} (t-1)^{\beta}}, \quad k = 1, \dots, t-1, \quad |c_{ik}(r, t)| \leq \frac{1}{2^{\alpha r} t^{\beta}}, \quad k = t. \quad (4.32)$$

Following the arguments in the proof of Theorem 3.1, we can get that for some  $\Gamma_7, \Gamma_7^* > 0$ ,

$$\begin{aligned} D_{rs}^{(2)} &\leq \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} P \left[ \max_{1 \leq n \leq 2^s} \left\| \sum_{t=2^s+1}^{2^s+n} \sum_{i=1}^{2^r} \sum_{k=1}^t c_{ik}(r, t) X_{ik} \right\| > \frac{\varepsilon}{4} \right] \\ &\leq \left( \frac{4}{\varepsilon} \right)^p \Gamma_7 \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \frac{[\log_2 2^{s+1}]^p}{2^{\alpha r p}} \sum_{i=1}^{2^r} \sum_{n=2^s+1}^{2^{s+1}} \sum_{k=1}^n |c_{ik}(r, t)|^p E \|X_{ik}\|^p \\ &\approx \left( \frac{4}{\varepsilon} \right)^p \Gamma_7 \sum_{r=u}^{\infty} \sum_{s=v}^{\infty} \frac{[s+1]^p}{2^{\alpha r p}} \sum_{i=1}^{2^r} \sum_{n=2^s+1}^{2^{s+1}} \left\{ \sum_{k=1}^{n-1} \frac{E \|X_{ik}\|^p}{n^{2\beta p}} + \frac{E \|X_{in}\|^p}{n^{\beta p}} \right\} \\ &\leq 2^{\beta p} \left( \frac{8}{\varepsilon} \right)^p \Gamma_7^* \left\{ \frac{1}{2^{\alpha u p}} \sum_{i=1}^{2^u} \sum_{k=1}^{2^v} \frac{E \|X_{ik}\|^p}{k^{\beta p}} + \frac{1}{2^{\alpha u p}} \sum_{i=1}^{2^u} \sum_{k=2^v+1}^{\infty} \frac{E \|X_{ik}\|^p}{k^{\beta p}} \right. \\ &\quad + \sum_{i=2^u+1}^{\infty} \sum_{k=1}^{2^v} \frac{E \|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} + \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E \|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} \\ &\quad \left. + \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E \|X_{ik}\|^p}{i^{\alpha p} k^{\beta p}} [1 + \log_2^p (k+1)] \right\}. \end{aligned} \quad (4.33)$$

The second term in (4.33) is obtained by the following basic calculation. Next, similar to the procedure of getting (4.24), we have

$$\begin{aligned} &\max_{2^r < m \leq 2^{r+1}} \max_{2^s < n \leq 2^{s+1}} \|\tau_{mn} - \tau_{m,2^s} - \tau_{2^r,n} + \tau_{2^r,2^s}\| \\ &\leq \max_{1 \leq m \leq 2^r} \max_{1 \leq n \leq 2^s} \left\| \sum_{t_1=2^r+1}^{2^r+m} \sum_{t_2=2^s+1}^{2^s+n} (\tau_{t_1,t_2} - \tau_{t_1-1,t_2} - \tau_{t_1,t_2-1} + \tau_{t_1-1,t_2-1}) \right\|, \end{aligned} \quad (4.34)$$

where

$$\begin{aligned} \tau_{t_1,t_2} - \tau_{t_1-1,t_2} - \tau_{t_1,t_2-1} + \tau_{t_1-1,t_2-1} &= \sum_{i=1}^{t_1} \sum_{k=1}^{t_2} d_{ik}(t_1, t_2) X_{ik}, \\ d_{ik}(t_1, t_2) &= \begin{cases} \left[ \frac{1}{t_1^\alpha} \left( 1 - \frac{i-1}{t_1} \right) - \frac{1}{(t_1-1)^\alpha} \left( 1 - \frac{i-1}{t_1-1} \right) \right] \left[ \frac{1}{t_2^\beta} - \frac{1}{(t_2-1)^\beta} \right], & k \leq n-1, \\ \left[ \frac{1}{t_1^\alpha} \left( 1 - \frac{i-1}{t_1} \right) - \frac{1}{(t_1-1)^\alpha} \left( 1 - \frac{i-1}{t_1-1} \right) \right] \frac{1}{t_2^\beta}, & k = n. \end{cases} \end{aligned} \quad (4.35)$$

Similarly

$$|d_{ik}(t_1, t_2)| \leq \frac{\beta}{t_1^{\alpha+1} t_2^{2\beta}}, \quad k \leq n, \quad |d_{ik}(t_1, t_2)| \leq \frac{1}{t_1^{\alpha+1} t_2^\beta}, \quad k = n. \quad (4.36)$$

Then, by the analogous approaches of the proof of Theorem 3.1, for some  $H > 0$ , we have

$$\begin{aligned} D_{rs}^{(3)} &\leq \left(\frac{4}{\varepsilon}\right)^p H \cdot \left\{ (\text{lower order terms}) + \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E|X_{ik}|^p}{i^{\alpha p} k^{\beta p}} \right. \\ &\quad \left. + \sum_{i=2^u+1}^{\infty} \sum_{k=2^v+1}^{\infty} \frac{E|X_{ik}|^p}{i^{\alpha p} k^{\beta p}} \cdot [1 + \log_2^p(k+1)] \right\}. \end{aligned} \quad (4.37)$$

Consequently, following the analogous arguments in the previous proof, we can conclude the desired result.  $\square$

*Proof of Theorem 3.3.* By  $\sum_{i=1}^{\infty} \sum_{k=1}^{\infty} (\tau_{ik}^p / i^{\alpha p} k^{\beta p}) = \infty$ , there exists an array  $\{\varepsilon_{ik}\}$  of nonincreasing positive numbers converging to 0 as  $\max\{i, k\} \rightarrow \infty$  such that  $\sum_{i=1}^{\infty} \sum_{k=1}^{\infty} (\tau_{ik}^p \varepsilon_{ik}^p / i^{\alpha p} k^{\beta p}) = \infty$ . Next, define an array of independent random variables  $\{W_{ik}\}$  with the following properties [8]:

$$P\left[W_{ik} = \frac{i^\alpha k^\beta}{\varepsilon_{ik}}\right] = \frac{\tau_{ik}^p \varepsilon_{ik}^p}{i^{\alpha p} k^{\beta p}}, \quad P[W_{ik} = 0] = 1 - \frac{\tau_{ik}^p \varepsilon_{ik}^p}{i^{\alpha p} k^{\beta p}}. \quad (4.38)$$

Then, it is easy to have that

$$E|W_{ik}|^p = \tau_{ik}^p, \quad 0 \leq E(W_{ik}) = \frac{\tau_{ik}^p \varepsilon_{ik}^{p-1}}{i^{\alpha(p-1)} k^{\beta(p-1)}} \leq \varepsilon_{ik}^{p-1}. \quad (4.39)$$

Now let  $X_{ik} = \tau_{ik} \cdot ((W_{ik} - E(W_{ik})) / \delta_{ik}) = X_{ik}^+ - X_{ik}^-$ , where  $\delta_{ik}^p = E|W_{ik} - E(W_{ik})|^p$ ,  $X_{ik}^+ = \tau_{ik} W_{ik} / \delta_{ik}$ , and  $X_{ik}^- = \tau_{ik} E(W_{ik}) / \delta_{ik}$ . Since  $W_{ik}$  and  $E(W_{ik})$  are all positive, and by the dominated convergence theorem for  $\delta_{ik}$ , we have

$$\frac{\tau_{ik}}{\delta_{ik}} \geq 1, \quad \lim_{i+k \rightarrow \infty} \frac{\tau_{ik}}{\delta_{ik}} = 1. \quad (4.40)$$

Consequently, we can obtain that  $X_{ik}^+ \geq \max\{0, W_{ik}\}$ ,  $i, k = 1, 2, 3, \dots$ , and  $|X_{ik}^-| \leq 2\varepsilon_{ik}^{p-1}$ , when  $i+k$  is sufficiently large. Then,

$$\begin{aligned}
& \left| \frac{1}{m^\alpha n^\beta} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) \left(1 - \frac{k-1}{n}\right) X_{ik} \right| \\
& \geq \left\| \left| \frac{1}{m^\alpha n^\beta} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) \left(1 - \frac{k-1}{n}\right) X_{ik}^+ \right| \right. \\
& \quad \left. - \left| \frac{1}{m^\alpha n^\beta} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) \left(1 - \frac{k-1}{n}\right) X_{ik}^- \right| \right\| \\
& \geq \frac{1}{m^\alpha n^\beta} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) \left(1 - \frac{k-1}{n}\right) X_{ik}^+ - \frac{2\varepsilon_{11}^{p-1}}{m^\alpha n^\beta} \left(\frac{m+1}{2}\right) \left(\frac{n+1}{2}\right) \quad (X_{ik}^+ \geq 0) \\
& \geq \frac{C}{m^\alpha n^\beta} \left\{ \sum_{i=[m/2]}^m \left(1 - \frac{i-1}{m}\right) \sum_{k=[n/2]}^n \left(1 - \frac{k-1}{n}\right) \right\} X_{[m/2], [n/2]}^+ - \frac{C_1 mn}{m^\alpha n^\beta}, \\
& \quad C, C_1 > 0, \quad (X_{ik}^+ \text{ are nondecreasing when } i+k \text{ is large enough}) \\
& \geq \frac{C}{m^\alpha n^\beta} \cdot \frac{m}{4} \cdot \frac{n}{4} \cdot X_{[m/2], [n/2]}^+ - \frac{C_1 mn}{m^\alpha n^\beta} \\
& \geq \frac{C}{m^\alpha n^\beta} \cdot \frac{m}{4} \cdot \frac{n}{4} \cdot \frac{\tau_{[m/2], [n/2]}}{\delta_{[m/2], [n/2]}} W_{[m/2], [n/2]} - \frac{C_1 mn}{m^\alpha n^\beta} \\
& \geq \frac{C}{m^\alpha n^\beta} \cdot \frac{m}{4} \cdot \frac{n}{4} \cdot \frac{[m/2]^\alpha [n/2]^\beta}{\varepsilon_{[m/2], [n/2]}} - \frac{C_1 mn}{m^\alpha n^\beta} \\
& \approx \frac{C \cdot m \cdot n}{64 \cdot \varepsilon_{[m/2], [n/2]}} - \frac{C_1 mn}{m^\alpha n^\beta} \approx \frac{C_2}{\varepsilon_{[m/2], [n/2]}} \cdot \frac{m \cdot n}{m^\alpha n^\beta}, \quad C_2 > 0,
\end{aligned} \tag{4.41}$$

where  $[\cdot]$  denotes the greatest integer function. Finally, employing the Borel-Cantelli lemma and the assumption conditions can yield that

$$\limsup_{m+n \rightarrow \infty} \left| \frac{1}{m^\alpha n^\beta} \sum_{i=1}^m \sum_{k=1}^n \left(1 - \frac{i-1}{m}\right) \left(1 - \frac{k-1}{n}\right) X_{ik} \right| = \infty, \quad \text{a.s.} \tag{4.42}$$

This completes the desired proof.  $\square$

*Proof of Theorem 3.4.* Here we define an array  $\{W_{ik}\}$  of independent random variables as follows [8]:

$$\begin{aligned}
P\left[ W_{ik} = \frac{i^\alpha k^\beta}{\varepsilon_{ik} \log_2(k+1)} \right] &= \frac{\tau_{ik}^p \varepsilon_{ik}^p}{i^{\alpha p} k^{\beta p}} \log_2^p(k+1), \\
P[W_{ik} = 0] &= 1 - \frac{\tau_{ik}^p \varepsilon_{ik}^p}{i^{\alpha p} k^{\beta p}} \log_2^p(k+1).
\end{aligned} \tag{4.43}$$

Hence, we also have that

$$E |W_{ik}|^p = \tau_{ik}^p, \quad 0 \leq E(W_{ik}) = \frac{\tau_{ik}^p \epsilon_{ik}^{p-1} \log_2^{p-1}(k+1)}{i^{\alpha(p-1)} k^{\beta(p-1)}} \leq \epsilon_{ik}^{p-1}. \quad (4.44)$$

Then, choosing  $\{X_{ik}\}$  and following the similar steps as in the proof of Theorem 3.3, we can obtain the desired results.  $\square$

### Acknowledgment

The author would like to thank the referee for his careful reading and valuable comments which make this paper more clear and readable.

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